

# CPSC 535

## Introduction to General State-Space Markov Chains

AD

April 2007

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- Starting from  $\mu \in M_1(E)$  and a sequence of kernels  $\{K_n; n \geq 1\}$  then

$$P_\mu(X_0 \in A_0, \dots, X_n \in A_n) = \int_{x_0 \in A_0} \cdots \int_{x_n \in A_n} \mu(dx_0) \prod_{i=1}^n K_i(x_{i-1}, dx_i).$$

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- We have

$$\mu(K(f)) = (\mu K)(f).$$



- Given  $K_1(x_1, dx_2)$  a Markov kernel from  $(E_1, \mathcal{E}_1)$  to  $(E_2, \mathcal{E}_2)$  and  $K_2(x_2, dx_3)$  a Markov kernel from  $(E_2, \mathcal{E}_2)$  to  $(E_3, \mathcal{E}_3)$  then we can define a new Markov kernel from  $(E_1, \mathcal{E}_1)$  to  $(E_3, \mathcal{E}_3)$

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$$K_n(x, A) = \int_{E^{n-1}} K(x, dx_1) K(x_1, dx_2) \cdots K(x_{n-1}, A)$$

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- In the MCMC context, we have typically  $X_0 \sim \mu \in M_1(E)$  and  $K$  an MCMC kernel of invariant distribution  $\pi$  and we want the measure

$$\mu K^n$$

to converge as fast as possible to  $\pi$ .

# Total Variation Norm

- We denote  $M(E)$  the space of bounded measures on  $(E, \mathcal{E})$  equipped with the norm

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- We can think of  $M_1(E)$  as

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- We can easily show that for any  $\mu \in M_1(E)$

$$\|\mu\| = \frac{1}{2}$$

and

$$\mu \in M_1(E) \Rightarrow \mu K \in M_1(E)$$

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- Clearly we have for any  $\mu_1, \mu_2 \in M_1(E)$  that  $\mu_1 - \mu_2 \in M_0(E)$  and

$$\|\mu_1 K - \mu_2 K\| \leq b(K) \|\mu_1 - \mu_2\|$$

so  $b(K)$  is a measure of the contraction induced by  $K$ .

- The number  $a(K) \in [0, 1]$  defined as follows

$$a(K) = \inf \left\{ \sum_{i=1}^m \min(K(x, A_i), K(y, A_i)) \right\}$$

where the infimum is taken over all points  $x, y \in E$ , the integers  $m \geq 1$  and the finite partitions  $\{A_i; 1 \leq i \leq m\}$  of  $E$ . It is called the **Dobrushin coefficient**.

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- We will show later on that

$$a(K) + b(K) = 1,$$

i.e. we want  $b(K)$  close to zero and  $a(K)$  close to one for fast mixing.

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- Moreover for any  $\mu_1, \mu_2 \in M_1(E)$  then

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- When the space is finite, then we have

$$\|\mu_1 - \mu_2\| = \frac{1}{2} \sum_{x \in E} |\mu_1(x) - \mu_2(x)|$$

and when  $\mu_1, \mu_2$  have densities  $f_1, f_2$  with respect to say  $\lambda$  then

$$\|\mu_1 - \mu_2\| = \frac{1}{2} \int_E |f_1(x) - f_2(x)| \lambda(dx).$$



- If  $\mu \in M_0(E)$  then for any  $A \in \mathcal{E}$  then

$$\mu(E) = 0 = \mu(A) + \mu(A^c) \text{ where } A^c = E - A.$$

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- Thus we have

$$\sup_{A \in \mathcal{E}} \mu(A) = -\sup_{A \in \mathcal{E}} \mu(A^c) = -\inf_{A \in \mathcal{E}} \mu(A)$$

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- Assume now that  $\mu_1, \mu_2 \in M_1(E)$  then clearly  $(\mu_1 - \mu_2) \in M_0(E)$  so it follows that

$$\begin{aligned} \|\mu_1 - \mu_2\| &= \sup_{A \in \mathcal{E}} (\mu_1(A) - \mu_2(A)) \\ &= \sup_{A \in \mathcal{E}} (\mu_2(A) - \mu_1(A)) \text{ (by symmetry)} \\ &= \sup_{A \in \mathcal{E}} \sup ((\mu_1(A) - \mu_2(A)), \mu_2(A) - \mu_1(A)) \\ &= \sup_{A \in \mathcal{E}} |\mu_1(A) - \mu_2(A)| \end{aligned}$$

- Now consider that  $E$  is a finite set and we denote

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- As for each  $A \in \mathcal{E}$  one has

$$\theta(A) = \frac{1}{2} (\theta(A) - \theta(A^c))$$

it follows that

$$\begin{aligned} \|\mu_1 - \mu_2\| &= \sup_{A \in \mathcal{E}} |\theta(A)| \leq \sup_{A \in \mathcal{E}} \frac{1}{2} \left( \sum_{x \in A} |\theta(x)| + \sum_{x \in A^c} |\theta(x)| \right) \\ &\leq \frac{1}{2} \sum_{x \in E} |\theta(x)| = \frac{1}{2} \sum_{x \in E} |\mu_1(x) - \mu_2(x)| \end{aligned}$$

- Moreover using once more that  $\theta(A) = \frac{1}{2}(\theta(A) - \theta(A^c))$ , then if we select

$$A = \{x \in E : \theta(x) \geq 0\}$$

then

$$\begin{aligned}\theta(A) &= \frac{1}{2} \left[ \sum_{x:\mu_1(x) \geq \mu_2(x)} \theta(x) - \sum_{x:\mu_1(x) \leq \mu_2(x)} \theta(x) \right] \\ &= \frac{1}{2} \left[ \sum_{x:\mu_1(x) \geq \mu_2(x)} |\theta(x)| + \sum_{x:\mu_1(x) \leq \mu_2(x)} |\theta(x)| \right] \\ &= \frac{1}{2} \sum_{x \in E} |\mu_1(x) - \mu_2(x)|.\end{aligned}$$

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- Hence it follows that

$$\|\mu_1 - \mu_2\| = \frac{1}{2} \sum_{x \in E} |\mu_1(x) - \mu_2(x)| = \sum_{x \in E: \mu_1(x) > \mu_2(x)} \mu_1(x) - \mu_2(x)$$

# Theorem (Dobrushin)

- For any Markov kernel  $K$  on  $E$ , the number  $b(K) = \sup_{\mu \in M_0(E)} \frac{\|\mu K\|}{\|\mu\|} \in [0, 1]$  can be written as

$$\begin{aligned} b(K) &= \sup_{\mu_1, \mu_2 \in M_1(E)} \|\mu_1 K - \mu_2 K\| / \|\mu_1 - \mu_2\| \\ &= \sup_{x, y \in E} \|K(x, \cdot) - K(y, \cdot)\| \\ &= 1 - a(K). \end{aligned}$$



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- Remark.** Showing that  $\sup_{x, y \in E} \|K(x, \cdot) - K(y, \cdot)\| = 1 - a(K)$  is equivalent to show that for any  $\mu_1, \mu_2 \in M_1(E)$  then

$$\|\mu_1 - \mu_2\| = 1 - \inf \left\{ \sum_{i=1}^m \min(\mu_1(A_i), \mu_2(A_i)) \right\}.$$

- **Corollary.** Assume there exists an integer  $p \geq 1$ ,  $\gamma \in M_1(E)$  and  $\varepsilon > 0$  such that for any  $(x, A) \in (E, \mathcal{E})$

$$K^p(x, A) \geq \varepsilon \gamma(A).$$

Thus  $K^p$  is a contracting operator on  $(M_0(E), \|\cdot\|)$  and for any  $\mu_1, \mu_2 \in M_1(E)$

$$\|\mu_1 K^p - \mu_2 K^p\| \leq (1 - \varepsilon) \|\mu_1 - \mu_2\|.$$

Moreover if  $K$  possesses an invariant measure  $\mu_\infty = \mu_\infty K$  then this one is unique and for any initial measure  $\mu \in M_1(E)$  then

$$\lim_{n \rightarrow \infty} \|\mu K^n - \mu_\infty\| = 0.$$

# Proof of Corollary

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- For any  $n \geq kp$  we have

$$\begin{aligned} \|\mu K^n - \mu_\infty K^n\| &\leq (1 - \varepsilon)^p \left\| \mu K^{n-kp} - \mu_\infty K^{n-kp} \right\| \\ &\leq (1 - \varepsilon)^p \|\mu - \mu_\infty\|. \end{aligned}$$

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- The invariant measure is obviously unique as if we had two then

$$\|\mu_\infty K^P - \nu_\infty K^P\| \leq (1 - \varepsilon) \|\mu_\infty - \nu_\infty\| \text{ (contraction)}$$

but

$$\|\mu_\infty K^P - \nu_\infty K^P\| = \|\mu_\infty - \nu_\infty\| \text{ (invariance).}$$

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- To prove that there is equality, consider a measure  $\mu \in M_0(E)$  and denote

$$A_\mu^+ = \{x \in E : \mu(x) \geq 0\}, \quad A_\mu^- = \{x \in E : \mu(x) < 0\}$$

then

$$\mu(E) = 0 \Rightarrow \mu(A_\mu^+) = -\mu(A_\mu^-).$$



- Hence we can rewrite  $\mu$  as a difference of probability measures up to a normalizing constant

$$\begin{aligned}\mu(B) &= \mu(A_\mu^+) \left( \frac{\mu(A_\mu^+ \cap B)}{\mu(A_\mu^+)} + \frac{\mu(A_\mu^- \cap B)}{\mu(A_\mu^+)} \right) \\ &= \mu(A_\mu^+) \left( \frac{\mu(A_\mu^+ \cap B)}{\mu(A_\mu^+)} - \frac{\mu(A_\mu^- \cap B)}{\mu(A_\mu^-)} \right)\end{aligned}$$

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- It follows that

$$b(K) = \sup_{\mu \in M_0(E)} \frac{\|\mu K\|}{\|\mu\|} = \sup_{\mu \in M_0(E)} \frac{\left\| \frac{\mu(A_\mu^+ \cap \cdot)}{\mu(A_\mu^+)} K - \frac{\mu(A_\mu^- \cap \cdot)}{\mu(A_\mu^-)} K \right\|}{\left\| \frac{\mu(A_\mu^+ \cap \cdot)}{\mu(A_\mu^+)} - \frac{\mu(A_\mu^- \cap \cdot)}{\mu(A_\mu^-)} \right\|}$$

and the result is proved.

- To show that  $b(K) = \sup_{x,y \in E} \|K(x, \cdot) - K(y, \cdot)\|$ , we can first show that

$$\begin{aligned} b(K) &= \sup_{\mu_1, \mu_2 \in M_1(E)} \|\mu_1 K - \mu_2 K\| / \|\mu_1 - \mu_2\| \\ &\geq \sup_{x,y \in E} \|\delta_x K - \delta_y K\| / \|\delta_x - \delta_y\| \\ &= \sup_{x,y \in E} \|\delta_x K - \delta_y K\| \end{aligned}$$

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 \end{aligned}$$

- To show the equality, remember that if  $\mu = \mu_1 - \mu_2$  where  $\mu_1, \mu_2 \in M_1(E)$  then

$$\|\mu\| = \frac{1}{2} \sum_{x \in E} |\mu(x)| = \sum_{x \in E: \mu(x) > 0} \mu(x) = - \sum_{x \in E: \mu(x) < 0} \mu(x).$$

- For any  $C \in \mathcal{E}$  and  $\mu = \mu_1 - \mu_2$ , we also have

$$\begin{aligned}
 \mu K(C) &= \sum_{\mu \geq 0} \mu(x) K(x, C) + \sum_{\mu < 0} \mu(x) K(x, C) \\
 &= \sum_{\mu \geq 0} \mu(x) K(x, C) - \sum_{\mu < 0} (-\mu(x)) K(x, C) \\
 &\leq \sum_{\mu \geq 0} \mu(x) K(x, C) - \left[ \inf_x K(x, C) \right] \sum_{\mu < 0} (-\mu(x)) \\
 &= \sum_{\mu \geq 0} \mu(y) \left[ K(y, C) - \inf_x K(x, C) \right] \\
 &\leq \sum_{\mu \geq 0} \mu(y) \left[ \sup_y K(y, C) - \inf_x K(x, C) \right] \\
 &= \|\mu\| \sup_{x,y} |K(y, C) - K(x, C)|
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 &= \|\mu\| \sup_{x,y} |K(y, C) - K(x, C)|
 \end{aligned}$$

- Now by taking the supremum on the  $C \in \mathcal{E}$ , the result follows.

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- *Proof.* We have

$$\begin{aligned} 2 \|\mu_1 - \mu_2\| &= \sum_{\mu_1 < \mu_2} (\mu_2(x) - \mu_1(x)) - \sum_{\mu_1 \geq \mu_2} (\mu_2(x) - \mu_1(x)) \\ &= 2 - \sum_{\mu_1 \geq \mu_2} \mu_2(x) - \sum_{\mu_1 < \mu_2} \mu_1(x) \\ &= 2 \left( 1 - \sum_{\mu_1 \geq \mu_2} \min(\mu_1(x), \mu_2(x)) - \sum_{\mu_1 < \mu_2} \min(\mu_1(x), \mu_2(x)) \right) \end{aligned}$$

thus

$$\|\mu_1 - \mu_2\| = 1 - \sum_{x \in E} \min(\mu_1(x), \mu_2(x))$$

- So it follows from this result that the measure  $\nu$  is defined by

$$\nu(x) = \min(\mu_1(x), \mu_2(x))$$

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- To prove the results in general measurable spaces then we need to use the Hahn-Jordan decomposition of the measure

$$\mu = \mu^+ - \mu^-.$$

# Application to Convergence of Simulated Annealing

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- We use a random walk Metropolis

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- We want to increase  $\beta$  to  $\infty$  as time increases as then  $\pi_{\beta}(x)$  concentrates itself on the set of global maxima of  $U(x)$ .

- Denoting  $\eta_0$  the initial distribution of  $X_0$ , then we have  
 $X_{n+1} | x_n \sim K_{\beta_n}(x_n, \cdot)$

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- The idea consists of using

$$\begin{aligned} \|\eta_{n+1} - \pi_{\beta_{n+1}}\| &= \|\eta_n K_{\beta_n} - \pi_{\beta_n} K_{\beta_n} + \pi_{\beta_n} - \pi_{\beta_{n+1}}\| \\ &\leq \underbrace{\|\eta_n K_{\beta_n} - \pi_{\beta_n} K_{\beta_n}\|}_{\text{mixing properties}} + \underbrace{\|\pi_{\beta_n} - \pi_{\beta_{n+1}}\|}_{\text{discrepancy successive targets}} \end{aligned}$$

- We have

$$\|\eta_n K_{\beta_n} - \pi_{\beta_n} K_{\beta_n}\| \leq \beta(K_{\beta_n}) \|\eta_n - \pi_{\beta_n}\|.$$

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- **Lemma.** For any  $\beta > 0$  and  $(x, y) \in E \times E$  then

$$K_\beta(x, y) \geq \exp(-\beta \operatorname{osc} U) q(x, y)$$

where

$$\operatorname{osc} U = \max_{x \in E} U(x) - \min_{x \in E} U(x).$$

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- *Proof.* Clearly we have

$$K_\beta(x, y) \geq \alpha_\beta(x, y) q(x, y)$$

where

$$\alpha_\beta(x, y) = \min(1, \exp(-\beta(U(y) - U(x)))) \geq \exp(-\beta \operatorname{osc} U).$$

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- **Proposition.** We have for any  $n > 0$

$$\begin{aligned} \left\| \eta_{n+1} - \pi_{\beta_{n+1}} \right\| &\leq \left\| \eta_n K_{\beta_n} - \pi_{\beta_n} K_{\beta_n} \right\| + \left\| \pi_{\beta_n} - \pi_{\beta_{n+1}} \right\| \\ &\leq (1 - \exp(-\beta_n \text{osc} U)) \left\| \eta_n - \pi_{\beta_n} \right\| + (\beta_{n+1} - \beta_n) \end{aligned}$$



- **Lemma.** Let  $I_n, a_n, b_n$  be three sequences positive numbers such that for  $n \geq 1$

$$I_n \leq (1 - a_n) I_{n-1} + b_n.$$

If

$$\lim_{n \rightarrow \infty} \frac{b_n}{a_n} = 0$$

and

$$\lim_{n \rightarrow \infty} \prod_{p=1}^n (1 - a_p) = 0$$

then

$$\lim_{n \rightarrow \infty} I_n = 0.$$

- *Proof.* For any  $\epsilon > 0$ ,  $\exists n(\epsilon) \geq 1$  such that for  $n \geq n(\epsilon)$

$$b_n \leq \epsilon a_n, \quad \prod_{p=1}^n (1 - a_p) \leq \epsilon.$$

Thus for  $n \geq n(\epsilon)$

$$\begin{aligned} l_n - \epsilon &\leq (1 - a_n) l_{n-1} - \epsilon (1 - a_n) \\ &= (1 - a_n) (l_{n-1} - \epsilon) \\ &\leq (l_0 - \epsilon) \prod_{p=1}^n (1 - a_p). \end{aligned}$$

It follows that

$$0 \leq l_n \leq \epsilon + \epsilon (l_0 + \epsilon) \leq \epsilon (1 + \epsilon + |l_0|).$$

The result follows.

- **Theorem.** Let  $\{X_n\}_{n \geq 0}$  be the simulated annealing scheme, then for any initial distribution  $\eta_0$  and  $\beta_n = \frac{\log(n+e)}{C}$ ,  $C > \text{osc}U$  then

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- **Theorem.** Let  $\{X_n\}_{n \geq 0}$  be the simulated annealing scheme, then for any initial distribution  $\eta_0$  and  $\beta_n = \frac{\log(n+e)}{C}$ ,  $C > \text{osc}U$  then

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- *Proof.* We have

$$\begin{aligned} \|\eta_{n+1} - \pi_{\beta_{n+1}}\| &\leq (1 - \exp(-\beta_n \text{osc}U)) \|\eta_n - \pi_{\beta_n}\| \\ &\quad + (\beta_{n+1} - \beta_n) \cdot \text{osc}U \end{aligned}$$

so by writing  $I_{n+1} = \|\eta_{n+1} - \pi_{\beta_{n+1}}\|$  then

$$I_{n+1} \leq (1 - a_{n+1}) I_n + b_{n+1}$$

where

$$a_{n+1} = \exp(-\beta_n \text{osc}U) = \frac{1}{(n+e)^{\frac{\text{osc}U}{C}}},$$

$$b_{n+1} = \frac{\text{osc}U}{C} \frac{1}{n+e}$$

- We have

$$\lim_{n \rightarrow \infty} \frac{b_n}{a_n} = \lim_{n \rightarrow \infty} \frac{1}{(n+e)^{1-\frac{\text{osc}U}{c}}} = 0$$

and

$$\begin{aligned} \prod_{p=1}^n (1-a_p) &\leq \exp\left(\sum_{p=1}^n \log(1-a_p)\right) \\ &\leq \exp\left(-\sum_{p=1}^n a_p\right) \xrightarrow{n \rightarrow \infty} 0 \end{aligned}$$